

Notes in RED are not part of solution.

Technical Exercise 1

(1) $\sum_{t=3}^6 X_t$ and $\sum_{j=3}^6 X_t$ are the same, while $\sum_{j=3}^6 X_j$ is not equal to either.

Observe that

$$\sum_{t=3}^6 X_t = X_3 + X_4 + X_5 + X_6$$

$$\sum_{j=3}^6 X_j = X_3 + X_4 + X_5 + X_6$$

$$\sum_{j=3}^6 X_t = X_t + X_t + X_t + X_t = 4X_t$$

j and t are really placeholders

$$(2) \sum_{i=1}^3 \sum_{j=2}^3 X_{ij} = \sum_{i=1}^3 (X_{i2} + X_{i3}) = \sum_{i=1}^3 X_{i2} + \sum_{i=1}^3 X_{i3} = X_{12} + X_{22} + X_{32} + X_{13} + X_{23} + X_{33}$$

↑
property (3)
in ExSet02

(3) Case of $n=2$:

$$\begin{aligned} \sum_{t=1}^2 (X_t - \bar{X})(Y_t - \bar{Y}) &= (X_1 - \bar{X})(Y_1 - \bar{Y}) + (X_2 - \bar{X})(Y_2 - \bar{Y}) \\ &= \left(X_1 - \frac{X_1 + X_2}{2}\right) \left(Y_1 - \frac{Y_1 + Y_2}{2}\right) + \left(X_2 - \frac{X_1 + X_2}{2}\right) \left(Y_2 - \frac{Y_1 + Y_2}{2}\right) \\ &= \frac{1}{2}(X_1 - X_2) \cdot \frac{1}{2}(Y_1 - Y_2) + \frac{1}{2}(X_2 - X_1) \frac{1}{2}(Y_2 - Y_1) \end{aligned}$$

$$= \frac{1}{4} (X_2 - X_1)(Y_2 - Y_1) + \frac{1}{4} (X_2 - X_1)(Y_2 - Y_1)$$

$$= \frac{1}{2} (X_2 - X_1)(Y_2 - Y_1)$$

$$\sum_{i=1}^2 \sum_{j=1}^2 (X_j - X_i)(Y_j - Y_i) = \sum_{i=1}^2 \left[(X_1 - X_i)(Y_1 - Y_i) + (X_2 - X_i)(Y_2 - Y_i) \right]$$

$$= (X_1 - X_1)(Y_1 - Y_1) + (X_2 - X_1)(Y_2 - Y_1) + (X_1 - X_2)(Y_1 - Y_2) \\ + (X_2 - X_2)(Y_2 - Y_2)$$

$$= 2(X_2 - X_1)(Y_2 - Y_1)$$

NOTE that I have a very big typo here in what I asked you to prove. I was too excited to share this result and I wanted to reduce the amount of work by excluding the denominator of the slope. Eventually, those constants $1/2$ and 2 would get canceled out. The correct equality is

$$\sum_{t=1}^n (X_t - \bar{X})(Y_t - \bar{Y}) = \frac{2}{n} \sum_{i=1}^n \sum_{j=1}^n (X_j - X_i)(Y_j - Y_i)$$

↳ The missing part.

Technical Exercise 2

$$(1) \text{ slope of regression of } Y \text{ on } X_1 = \frac{\sum_{t=1}^n (X_{1t} - \bar{X}_1)(Y_t - \bar{Y})}{\sum_{t=1}^n (X_{1t} - \bar{X}_1)^2}$$

$$(2) \text{ slope of regression of } Z \text{ on } W = \frac{\sum_{t=1}^n (W_t - \bar{W})(Z_t - \bar{Z})}{\sum_{t=1}^n (W_t - \bar{W})^2}$$

(3) observe that

$$\begin{aligned} \sum_{t=1}^n (W_t - \bar{W})(Z_t - \bar{Z}) &= \sum_{t=1}^n (aX_{1t} + b - (a\bar{X}_1 + b))(cY_t + d - (c\bar{Y} + d)) \\ &= \sum_{t=1}^n a(X_{1t} - \bar{X}_1)c(Y_t - \bar{Y}) \\ &= ac \sum_{t=1}^n (X_{1t} - \bar{X}_1)(Y_t - \bar{Y}) \end{aligned}$$

Next,

$$\sum_{t=1}^n (W_t - \bar{W})^2 = \sum_{t=1}^n (aX_{1t} + b - (a\bar{X}_1 + b))^2 = \sum_{t=1}^n a^2(X_{1t} - \bar{X}_1)^2 = a^2 \sum_{t=1}^n (X_{1t} - \bar{X}_1)^2$$

Therefore,

$$\text{slope of regression of } Z \text{ on } W = \frac{ac \sum_{t=1}^n (X_{1t} - \bar{X}_1)(Y_t - \bar{Y})}{a^2 \sum_{t=1}^n (X_{1t} - \bar{X}_1)^2} = \frac{c}{a} \cdot \text{slope of regression of } Y \text{ on } X_1$$

Technical Exercise 3

① Observe that

$$\begin{aligned} \sum_{t=1}^n (Y_t - \hat{\beta}_1 X_{1t})^2 &= \sum_{t=1}^n (Y_t^2 - 2\hat{\beta}_1 Y_t X_{1t} + \hat{\beta}_1^2 X_{1t}^2) \\ &= \sum_{t=1}^n Y_t^2 - 2\hat{\beta}_1 \sum_{t=1}^n Y_t X_{1t} + \hat{\beta}_1^2 \sum_{t=1}^n X_{1t}^2 \end{aligned}$$

The first-order condition is

$$\frac{d}{d\hat{\beta}_1} \left[\sum_{t=1}^n Y_t^2 - 2\hat{\beta}_1 \sum_{t=1}^n Y_t X_{1t} + \hat{\beta}_1^2 \sum_{t=1}^n X_{1t}^2 \right] = 0$$

$$-2 \sum_{t=1}^n Y_t X_{1t} + 2\hat{\beta}_1 \sum_{t=1}^n X_{1t}^2 = 0$$

$$\hat{\beta}_1 = \frac{\sum_{t=1}^n X_{1t} Y_t}{\sum_{t=1}^n X_{1t}^2}$$

The second-order condition is given by

$$\frac{d}{d\hat{\beta}_1} \left[-2 \sum_{t=1}^n Y_t X_{1t} + 2\hat{\beta}_1 \sum_{t=1}^n X_{1t}^2 \right] = 2 \sum_{t=1}^n X_{1t}^2 \quad \hat{\beta}_1 \text{ will be optimal}$$

(minimize sum of squared residuals) if $\sum_{t=1}^n X_{1t}^2 > 0$ which is always true
(Sum of positive numbers is positive)

with one exception. The exception happens when all observations of X_1 are all equal to zero meaning $X_{1t} = 0$ for all $t = 1, \dots, n$. But this would never happen because $\hat{\beta}_1$ will not even exist in this situation.

② The residuals are $e_t = Y_t - \hat{\beta}_1 X_{1t} = Y_t - \left(\frac{\sum_{t=1}^n X_{1t} Y_t}{\sum_{t=1}^n X_{1t}^2} \right) X_{1t}$. → really an expression which no longer depends on t

Taking averages, we have

$$\begin{aligned} \frac{1}{n} \sum_{t=1}^n e_t &= \frac{1}{n} \sum_{t=1}^n Y_t - \left(\frac{\sum_{t=1}^n Y_t X_{1t}}{\sum_{t=1}^n X_{1t}^2} \right) \frac{1}{n} \sum_{t=1}^n X_{1t} \\ &= \bar{Y} - \left(\frac{\sum_{t=1}^n Y_t X_{1t}}{\sum_{t=1}^n X_{1t}^2} \right) \bar{X}_1 \end{aligned}$$

There is no way to guarantee that the average of the residuals is equal to zero. It will only happen when \bar{Y} and \bar{X}_1 are both equal to zero.